

Is there gender discrimination in wage? Using dummy variables and interactions

Source: These are data from the 2006 Current Population Survey. 2000 working adults

wage	float	%9.0g	Average hourly earnings (in \$)
educ	byte	%8.0g	Years of education
exper	byte	%8.0g	Potential years of experience
female	byte	%8.0g	Female
union	byte	%8.0g	Union member
cateduc	float	%9.0g	Educ: incomplete high, high sch., some college

1. Estimating difference in means between male and female:

Variable	Obs	Mean	Std. Dev.	Min	Max
female	1033	1	0	1	1
wage	1033	16.12258	9.715608	2.125	72.125

Variable	Obs	Mean	Std. Dev.	Min	Max
female	967	0	0	0	0
wage	967	20.72326	12.71402	.7	82.42857

Test?

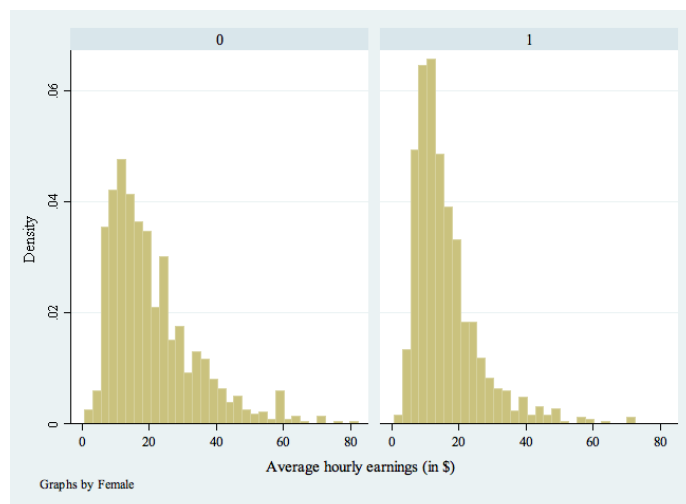
```
. ttest wage, by(female);
```

Two-sample t test with equal variances

Group	Obs	Mean	Std. Err.	Std. Dev.	[95% Conf. Interval]	
0	967	20.72326	.4088552	12.71402	19.92091	21.52561
1	1033	16.12258	.3022872	9.715608	15.52942	16.71575
combined	2000	18.34701	.2570348	11.49495	17.84293	18.85109
diff		4.600677	.5040778		3.612104	5.58925
diff = mean(0) - mean(1)						t = 9.1269
Ho: diff = 0				degrees of freedom =		1998
Ha: diff < 0		Ha: diff != 0		Ha: diff > 0		
Pr(T < t) = 1.0000		Pr(T > t) = 0.0000		Pr(T > t) = 0.0000		

2. Whole distribution of wage (hourly earnings in dollars)

```
. histogram wage, by(female)
```



3. Discrimination, even after controlling for difference in characteristics: Additive female effect.

```
. reg wage female union educ exper;
```

Source	SS	df	MS	Number of obs =	2000
Model	66564.2059	4	16641.0515	F(4, 1995) =	168.04
Residual	197571.272	1995	99.0332189	Prob > F =	0.0000
Total	264135.478	1999	132.133806	R-squared =	0.2520
				Adj R-squared =	0.2505
				Root MSE =	9.9515

wage	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
female	-5.178971	.4459606	-11.61	0.000	-6.053568 -4.304374
union	2.250108	.6475175	3.47	0.001	.9802265 3.519989
educ	2.225234	.107038	20.79	0.000	2.015316 2.435152
exper	.1759976	.0175156	10.05	0.000	.1416468 .2103485
_cons	-13.25096	1.521768	-8.71	0.000	-16.23538 -10.26654

Test? Interpretation?

Estimating difference in means with a simple regression, not controlling for characteristics:

```
. reg wage female;
```

Source	SS	df	MS	Number of obs =	2000
Model	10571.589	1	10571.589	F(1, 1998) =	83.30
Residual	253563.889	1998	126.908853	Prob > F =	0.0000
Total	264135.478	1999	132.133806	R-squared =	0.0400
				Adj R-squared =	0.0395
				Root MSE =	11.265

wage	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
female	-4.600677	.5040778	-9.13	0.000	-5.58925 -3.612104
_cons	20.72326	.3622703	57.20	0.000	20.01279 21.43373

4. Do females have differential return to some characteristics?

Is there a differential effect of union on women and men's wage: interaction between dummy variables

```
. gen femunion=female*union;
. reg wage female union femunion educ exper;
```

Source	SS	df	MS	Number of obs =	2000
Model	66586.3651	5	13317.273	F(5, 1994) =	134.42
Residual	197549.112	1994	99.0717715	Prob > F =	0.0000
Total	264135.478	1999	132.133806	R-squared =	0.2521
				Adj R-squared =	0.2502
				Root MSE =	9.9535

wage	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
female	-5.094864	.4801934	-10.61	0.000	-6.036597 -4.15313
union	2.577524	.9480136	2.72	0.007	.7183235 4.436725
femunion	-.612306	1.29469	-0.47	0.636	-3.151394 1.926782
educ	2.228979	.1073513	20.76	0.000	2.018447 2.439512
exper	.1756898	.0175311	10.02	0.000	.1413085 .2100711
_cons	-13.33871	1.533331	-8.70	0.000	-16.3458 -10.33161

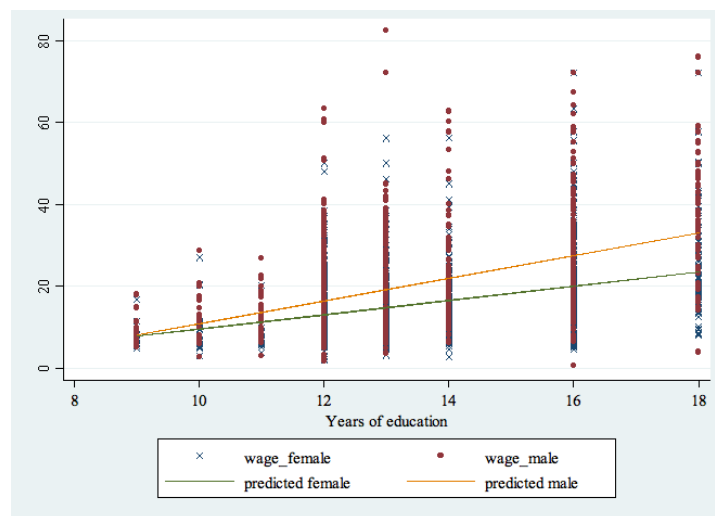
Is there a differential return to education for male and female? Interaction between dummy and continuous variables

```
. g femeduc=female*educ
. reg wage female educ femeduc
```

Source	SS	df	MS	Number of obs =	2000
Model	56945.4372	3	18981.8124	F(3, 1996) =	182.86
Residual	207190.04	1996	103.802625	Prob > F =	0.0000
Total	264135.478	1999	132.133806	R-squared =	0.2156
				Adj R-squared =	0.2144
				Root MSE =	10.188

wage	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
female	8.838706	3.013838	2.93	0.003	2.928108 14.7493
educ	2.772576	.1560434	17.77	0.000	2.466551 3.078601
femeduc	-1.019981	.2186047	-4.67	0.000	-1.448698 -.5912633
_cons	-16.7825	2.136137	-7.86	0.000	-20.97179 -12.59321

Female effect on wage = $(8.8 - 1.02 \text{ educ})$
 Education effect on wage = $(2.77 - 1.02 \text{ female})$



```
* graph;
qui reg wage female educ femeduc ;
predict wagehat;
gen wage_female=wage if female==1;
gen wage_male=wage if female==0;
gen trfem=wagehat if female==1;
gen trmale=wagehat if female==0;
label variable trfem "predicted female";
label variable trmale "predicted male";
twoway scatter wage_female wage_male trfem trmale educ, ms(X o i i) c(i i l l);
```

General case of interaction terms

Does the marginal effect of experience depend on education?

```
. gen expeduc=exper*educ
. reg wage female educ exper expeduc;
```

Source	SS	df	MS	Number of obs = 2000		
Model	65400.0183	4	16350.0046	F(4, 1995) = 164.13		
Residual	198735.459	1995	99.6167715	Prob > F = 0.0000		
Total	264135.478	1999	132.133806	R-squared = 0.2476		
				Adj R-squared = 0.2461		
				Root MSE = 9.9808		

wage	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
female	-5.145376	.4478427	-11.49	0.000	-6.023664	-4.267087
educ	2.158549	.1985612	10.87	0.000	1.76914	2.547958
exper	.1162292	.1165983	1.00	0.319	-.112438	.3448965
expeduc	.0049071	.0087011	0.56	0.573	-.0121571	.0219713
_cons	-12.19337	2.675967	-4.56	0.000	-17.44135	-6.945383

5. Use of ordinal variables

If education is given in 3 levels: cateduc =1 for high school dropout, =2 for high school, and =3 for some college education.

```
. gen cateduc1 = cateduc ==1
. gen cateduc2 = cateduc ==2
. gen cateduc3 = cateduc ==3
. reg wage female cateduc2 cateduc3 exper;
```

Source	SS	df	MS	Number of obs = 2000		
Model	43948.4454	4	10987.1113	F(4, 1995) = 99.55		
Residual	220187.032	1995	110.36944	Prob > F = 0.0000		
Total	264135.478	1999	132.133806	R-squared = 0.1664		
				Adj R-squared = 0.1647		
				Root MSE = 10.506		

wage	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
female	-5.144792	.4713192	-10.92	0.000	-6.069121	-4.220462
cateduc2	3.851033	1.011044	3.81	0.000	1.868221	5.833845
cateduc3	9.865773	.9566261	10.31	0.000	7.989682	11.74186
exper	.1773909	.0187487	9.46	0.000	.1406219	.2141599
_cons	9.996611	.9495125	10.53	0.000	8.134471	11.85875

Would it make sense to treat cateduc as if it was a real number?

```
. reg wage female cateduc exper;
```

...

wage	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
female	-5.154772	.4714936	-10.93	0.000	-6.079443	-4.230101
cateduc	5.417826	.3805958	14.24	0.000	4.671419	6.164233

...

6. Are the wage equations for male and female the same?

```
. reg wage educ exper;
```

Source	SS	df	MS	Number of obs =	2000
Model	52115.2387	2	26057.6194	F(2, 1997) =	245.43
Residual	212020.239	1997	106.169373	Prob > F =	0.0000
				R-squared =	0.1973
				Adj R-squared =	0.1965
Total	264135.478	1999	132.133806	Root MSE =	10.304

wage	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
educ	2.192429	.1103889	19.86	0.000	1.97594 2.408919
exper	.1770016	.0180643	9.80	0.000	.1415748 .2124285
_cons	-15.18658	1.566867	-9.69	0.000	-18.25944 -12.11371

```
. reg wage educ exper if female==1;
```

Source	SS	df	MS	Number of obs =	1033
Model	16027.8434	2	8013.92168	F(2, 1030) =	101.42
Residual	81385.7713	1030	79.015312	Prob > F =	0.0000
				R-squared =	0.1645
				Adj R-squared =	0.1629
Total	97413.6147	1032	94.3930375	Root MSE =	8.8891

wage	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
educ	1.795799	.1337998	13.42	0.000	1.533248 2.05835
exper	.1210147	.021846	5.54	0.000	.0781469 .1638824
_cons	-11.05997	1.938354	-5.71	0.000	-14.86355 -7.256399

```
. reg wage educ exper if female==0;
```

Source	SS	df	MS	Number of obs =	967
Model	41545.5802	2	20772.7901	F(2, 964) =	174.73
Residual	114604.694	964	118.884537	Prob > F =	0.0000
				R-squared =	0.2661
				Adj R-squared =	0.2645
Total	156150.274	966	161.646246	Root MSE =	10.903

wage	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
educ	2.6853	.1673038	16.05	0.000	2.356978 3.013621
exper	.2352962	.0273877	8.59	0.000	.1815499 .2890426
_cons	-20.38104	2.324115	-8.77	0.000	-24.94195 -15.82013

```
. g femexper=female*exper
```

```
. reg wage educ exper female femeduc femexper;
```

Source	SS	df	MS	Number of obs =	2000
Model	68145.0125	5	13629.0025	F(5, 1994) =	138.66
Residual	195990.465	1994	98.2901028	Prob > F =	0.0000
				R-squared =	0.2580
				Adj R-squared =	0.2561
Total	264135.478	1999	132.133806	Root MSE =	9.9141

wage	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
educ	2.6853	.1521241	17.65	0.000	2.386961 2.983638
exper	.2352962	.0249028	9.45	0.000	.1864581 .2841344
female	9.321065	3.023169	3.08	0.002	3.392165 15.24996
femeduc	-.8895005	.213099	-4.17	0.000	-1.307421 -.4715804
femexper	-.1142815	.0348398	-3.28	0.001	-.1826078 -.0459553
_cons	-20.38104	2.113245	-9.64	0.000	-24.52544 -16.23664

```
. test fmeduc femexper;
( 1) fmeduc = 0
( 2) femexper = 0

F( 2, 1994) = 14.12
Prob > F = 0.0000
```

```
. test female fmeduc femexper;
( 1) female = 0
( 2) fmeduc = 0
( 3) femexper = 0

F( 3, 1994) = 54.36
Prob > F = 0.0000
```

The New York Times

December 12, 2002, Thursday, Late Edition - Final

Economic Scene; Sticks and stones can break bones, but the wrong name can make a job hard to find. By

Alan B. Krueger

WHAT'S in a name? Evidently plenty if you are looking for a job.

To test whether employers discriminate against black job applicants, Marianne Bertrand of the University of Chicago and Sendhil Mullainathan of M.I.T. conducted an unusual experiment. They selected 1,300 help-wanted ads from newspapers in Boston and Chicago and submitted multiple resumes from phantom job seekers. The researchers randomly assigned the first names on the resumes, choosing from one set that is particularly common among blacks and from another that is common among whites.

So Kristen and Tamika, and Brad and Tyrone, applied for jobs from the same pool of want ads and had equivalent resumes. Nine names were selected to represent each category: black women, white women, black men and white men. Last names common to the racial group were also assigned. Four resumes were typically submitted for each job opening, drawn from a reservoir of 160. Nearly 5,000 applications were submitted from mid-2001 to mid-2002. Professors Bertrand and Mullainathan kept track of which candidates were invited for job interviews.

No single employer was sent two identical resumes, and the names on the resumes were randomly assigned, so applicants with black- and white-sounding names applied for the same set of jobs with the same set of resumes.

Apart from their names, applicants had the same experience, education and skills, so employers had no reason to distinguish among them.

The results are disturbing. Applicants with white-sounding names were 50 percent more likely to be called for interviews than were those with black-sounding names. Interviews were requested for 10.1 percent of applicants with white-sounding names and only 6.7 percent of those with black-sounding names.

Within racial groups, applications with men's or women's names were equally likely to result in calls for interviews, providing little evidence of discrimination based on sex in these entry-level jobs.

Their most alarming finding is that the likelihood of being called for an interview rises sharply with an applicant's credentials -- like experience and honors -- for those with white-sounding names, but much less for those with black-sounding names. A grave concern is that this phenomenon may be damping the incentives for blacks to acquire job skills, producing a self-fulfilling prophecy that perpetuates prejudice and misallocates resources.

(Source: "Are Emily and Brendan More Employable than Lakisha and Jamal? A Field Experiment on Labor Market Discrimination" *The American Economic Review*, 2004)