



### 3. Discrimination, even after controlling for difference in characteristics: Additive female effect.

```
. tabstat wage educ exper union service profocc , by(female)
```

Summary statistics: mean  
by categories of: female (Female)

female	wage	educ	exper	union	services	profocc
0	20.72326	13.5274	20.3061	.1323681	.1323681	.1664943
1	16.12258	13.73185	20.84608	.1452081	.1703775	.2545983
Total	18.34701	13.633	20.585	.139	.152	.212

And yes, the wages are lower in the service sector (average of \$12.9 compared to \$19.3) and higher for professional occupation (average of \$22.9 compared to \$17.1)

```
Use: tabstat wage, by(service)
```

Note that the female wages are lower even within the same sector, or within the same professional category

```
. table female services, c(mean wage)
```

Female	Works in service sector	
	0	1
0	21.62085	14.83982
1	17.0615	11.55071

```
. reg wage female educ exper union service profocc
```

Source	SS	df	MS	Number of obs = 2000		
Model	69360.4741	6	11560.079	F( 6, 1993) = 118.29		
Residual	194775.003	1993	97.7295551	Prob > F = 0.0000		
Total	264135.478	1999	132.133806	R-squared = 0.2626		
				Adj R-squared = 0.2604		
				Root MSE = 9.8858		

wage	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
female	-5.161436	.4466179	-11.56	0.000	-6.037323	-4.285549
educ	2.000723	.1182173	16.92	0.000	1.76888	2.232565
exper	.1711896	.0175183	9.77	0.000	.1368336	.2055457
union	2.229349	.6460632	3.45	0.001	.962319	3.496379
services	-2.729558	.6422974	-4.25	0.000	-3.989203	-1.469914
profocc	1.531844	.6118991	2.50	0.012	.3318149	2.731873
_cons	-10.00724	1.645929	-6.08	0.000	-13.23516	-6.779315

**Test? Interpretation?**

### Estimating difference in means with a simple regression, not controlling for characteristics:

```
. reg wage female;
```

Source	SS	df	MS	Number of obs = 2000		
Model	10571.589	1	10571.589	F( 1, 1998) = 83.30		
Residual	253563.889	1998	126.908853	Prob > F = 0.0000		
Total	264135.478	1999	132.133806	R-squared = 0.0400		
				Adj R-squared = 0.0395		
				Root MSE = 11.265		

wage	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
female	-4.600677	.5040778	-9.13	0.000	-5.58925	-3.612104
_cons	20.72326	.3622703	57.20	0.000	20.01279	21.43373

#### 4. Do females have differential return to some characteristics?

Is there a differential return to education for male and female? Interaction between dummy and continuous variables

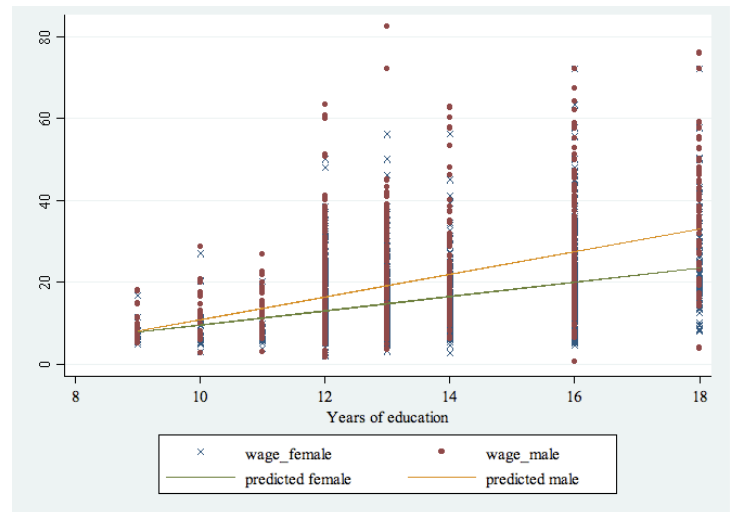
```
. g femeduc=female*educ
. reg wage female educ femeduc
```

Source	SS	df	MS	Number of obs =	2000
Model	56945.4372	3	18981.8124	F( 3, 1996) =	182.86
Residual	207190.04	1996	103.802625	Prob > F =	0.0000
Total	264135.478	1999	132.133806	R-squared =	0.2156
				Adj R-squared =	0.2144
				Root MSE =	10.188

wage	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
female	8.838706	3.013838	2.93	0.003	2.928108 14.7493
educ	2.772576	.1560434	17.77	0.000	2.466551 3.078601
femeduc	-1.019981	.2186047	-4.67	0.000	-1.448698 -.5912633
_cons	-16.7825	2.136137	-7.86	0.000	-20.97179 -12.59321

Female effect on wage = (8.8 - 1.02 educ)  
 Education effect on wage = (2.77 - 1.02 female)

```
* graph;
qui reg wage female educ femeduc ;
predict wagehat;
gen wage_female=wage if female==1;
gen wage_male=wage if female==0;
gen trfem=wagehat if female==1;
gen trmale=wagehat if female==0;
label variable trfem "predicted female";
label variable trmale "predicted male";
twoway scatter wage_female wage_male trfem
trmale educ, ms(x o i i) c(i i l l);
```



Is there a differential effect of union on women and men's wage: interaction between dummy variables

```
. gen femunion=female*union;
. reg wage female union femunion educ exper;
```

Source	SS	df	MS	Number of obs =	2000
Model	66586.3651	5	13317.273	F( 5, 1994) =	134.42
Residual	197549.112	1994	99.0717715	Prob > F =	0.0000
Total	264135.478	1999	132.133806	R-squared =	0.2521
				Adj R-squared =	0.2502
				Root MSE =	9.9535

wage	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
female	-5.094864	.4801934	-10.61	0.000	-6.036597 -4.15313
union	2.577524	.9480136	2.72	0.007	.7183235 4.436725
femunion	-.612306	1.29469	-0.47	0.636	-3.151394 1.926782
educ	2.228979	.1073513	20.76	0.000	2.018447 2.439512
exper	.1756898	.0175311	10.02	0.000	.1413085 .2100711
_cons	-13.33871	1.533331	-8.70	0.000	-16.3458 -10.33161

**General case of interaction terms**  
**Does the marginal effect of experience depend on education?**

```
. gen expeduc=exper*educ
. reg wage female educ exper expeduc;
```

Source	SS	df	MS	Number of obs =	2000
Model	65400.0183	4	16350.0046	F( 4, 1995) =	164.13
Residual	198735.459	1995	99.6167715	Prob > F =	0.0000
Total	264135.478	1999	132.133806	R-squared =	0.2476
				Adj R-squared =	0.2461
				Root MSE =	9.9808

wage	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
female	-5.145376	.4478427	-11.49	0.000	-6.023664 -4.267087
educ	2.158549	.1985612	10.87	0.000	1.76914 2.547958
exper	.1162292	.1165983	1.00	0.319	-.112438 .3448965
expeduc	.0049071	.0087011	0.56	0.573	-.0121571 .0219713
_cons	-12.19337	2.675967	-4.56	0.000	-17.44135 -6.945383

**5. Use of ordinal variables**

If education is given in 3 levels: cateduc =1 for high school dropout, =2 for high school, and =3 for some college education.

```
. gen HSDrop = cateduc ==1
. gen HS = cateduc ==2
. gen Col = cateduc ==3
. reg wage female HS Col exper
```

Source	SS	df	MS	Number of obs =	2000
Model	43948.4454	4	10987.1113	F( 4, 1995) =	99.55
Residual	220187.032	1995	110.36944	Prob > F =	0.0000
Total	264135.478	1999	132.133806	R-squared =	0.1664
				Adj R-squared =	0.1647
				Root MSE =	10.506

wage	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
female	-5.144792	.4713192	-10.92	0.000	-6.069121 -4.220462
HS	3.851033	1.011044	3.81	0.000	1.868221 5.833845
Col	9.865773	.9566261	10.31	0.000	7.989682 11.74186
exper	.1773909	.0187487	9.46	0.000	.1406219 .2141599
_cons	9.996611	.9495125	10.53	0.000	8.134471 11.85875

**Would it make sense to treat cateduc as if it was a real number?**

```
. reg wage female cateduc exper;
```

.....

wage	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
female	-5.154772	.4714936	-10.93	0.000	-6.079443 -4.230101
cateduc	5.417826	.3805958	14.24	0.000	4.671419 6.164233

.....

## 6. Are the wage equations for male and female the same?

```
. reg wage educ exper;
```

Source	SS	df	MS	Number of obs =	2000
Model	52115.2387	2	26057.6194	F( 2, 1997) =	245.43
Residual	212020.239	1997	106.169373	Prob > F =	0.0000
				R-squared =	0.1973
				Adj R-squared =	0.1965
Total	264135.478	1999	132.133806	Root MSE =	10.304

wage	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
educ	2.192429	.1103889	19.86	0.000	1.97594	2.408919
exper	.1770016	.0180643	9.80	0.000	.1415748	.2124285
_cons	-15.18658	1.566867	-9.69	0.000	-18.25944	-12.11371

```
. reg wage educ exper if female==1;
```

Source	SS	df	MS	Number of obs =	1033
Model	16027.8434	2	8013.92168	F( 2, 1030) =	101.42
Residual	81385.7713	1030	79.015312	Prob > F =	0.0000
				R-squared =	0.1645
				Adj R-squared =	0.1629
Total	97413.6147	1032	94.3930375	Root MSE =	8.8891

wage	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
educ	1.795799	.1337998	13.42	0.000	1.533248	2.05835
exper	.1210147	.021846	5.54	0.000	.0781469	.1638824
_cons	-11.05997	1.938354	-5.71	0.000	-14.86355	-7.256399

```
. reg wage educ exper if female==0;
```

Source	SS	df	MS	Number of obs =	967
Model	41545.5802	2	20772.7901	F( 2, 964) =	174.73
Residual	114604.694	964	118.884537	Prob > F =	0.0000
				R-squared =	0.2661
				Adj R-squared =	0.2645
Total	156150.274	966	161.646246	Root MSE =	10.903

wage	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
educ	2.6853	.1673038	16.05	0.000	2.356978	3.013621
exper	.2352962	.0273877	8.59	0.000	.1815499	.2890426
_cons	-20.38104	2.324115	-8.77	0.000	-24.94195	-15.82013

```
. g femexper=female*exper
```

```
. reg wage educ exper female femeduc femexper;
```

Source	SS	df	MS	Number of obs =	2000
Model	68145.0125	5	13629.0025	F( 5, 1994) =	138.66
Residual	195990.465	1994	98.2901028	Prob > F =	0.0000
				R-squared =	0.2580
				Adj R-squared =	0.2561
Total	264135.478	1999	132.133806	Root MSE =	9.9141

wage	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
educ	2.6853	.1521241	17.65	0.000	2.386961	2.983638
exper	.2352962	.0249028	9.45	0.000	.1864581	.2841344
female	9.321065	3.023169	3.08	0.002	3.392165	15.24996
femeduc	-.8895005	.213099	-4.17	0.000	-1.307421	-.4715804
femexper	-.1142815	.0348398	-3.28	0.001	-.1826078	-.0459553
_cons	-20.38104	2.113245	-9.64	0.000	-24.52544	-16.23664

```

. test femeduc femexper;
( 1) femeduc = 0
( 2) femexper = 0
F( 2, 1994) = 14.12
Prob > F = 0.0000

. test female femeduc femexper;
( 1) female = 0
( 2) femeduc = 0
( 3) femexper = 0
F( 3, 1994) = 54.36
Prob > F = 0.0000

```

## 7. Pooled Cross Sections – Looking at the evolution of gender wage gap over time.

**Source:** Compiled from the 1996 and 2006 Current Population Surveys

```

union          float %8.0g          =1 if respondent is union member
wage           float %9.0g          average hourly earnings (in $)
female        float %9.0g          1=female, 0=male
educ          float %9.0g          years of education
exper         float %9.0g          years potential experience
year          float %9.0g          1996 or 2006

```

Variable	Obs	Mean	Std. Dev.	Min	Max
union	4000	.98975	.9251264	0	2
wage	4000	15.59488	10.43896	.05	144.25
female	4000	.50125	.5000609	0	1
educ	4000	13.5125	2.113637	9	18
exper	4000	19.7475	12.54303	0	64
year	4000	2001	5.000625	1996	2006

### a. To compare wages across different years, need to use real wages

Use the Consumer Price Index (CPI) 1996: 156.9, 2006: 201.6, i.e., inflation has been 28.5% over 10 years, which is in average 2.53% per year

```

. g rwage=wage
. replace rwage=wage*201.6/156.9) if year==1996
. label variable rwage "average real hourly earnings (in 2006$)"
. bysort year: sum wage rwage;

```

-> year = 1996

Variable	Obs	Mean	Std. Dev.	Min	Max
wage	2000	12.69577	7.907586	.125	62.5
rwage	2000	16.31273	10.16042	.1606119	80.30593

-> year = 2006

Variable	Obs	Mean	Std. Dev.	Min	Max
wage	2000	18.49399	11.77498	.05	144.25
rwage	2000	18.49399	11.77498	.05	144.25

```
. gen yr2006=(year==2006);
```

```
. reg rwage educ exper union female yr2006;
```

Source	SS	df	MS	Number of obs =	4000
Model	131177.324	5	26235.4647	F( 5, 3994) =	293.43
Residual	357107.093	3994	89.4108896	Prob > F =	0.0000
				R-squared =	0.2686
				Adj R-squared =	0.2677
Total	488284.417	3999	122.10163	Root MSE =	9.4557

rwage	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
-------	-------	-----------	---	------	----------------------

educ	2.327535	.071202	32.69	0.000	2.187939	2.467131
exper	.1504879	.0119518	12.59	0.000	.1270557	.1739202
union	-.7340047	.4264605	-1.72	0.085	-1.570105	.102096
female	-4.748954	.299762	-15.84	0.000	-5.336655	-4.161253
yr2006	-.2317524	.7904708	-0.29	0.769	-1.781516	1.318012
_cons	-13.79644	1.249616	-11.04	0.000	-16.24639	-11.3465

## b. What happened to the gender gap?

```
. gen yr2006female=yr2006*female;
. reg rwage educ exper union female yr2006 yr2006female;
```

Source	SS	df	MS	Number of obs =	4000
Model	131623.93	6	21937.3217	F( 6, 3993) =	245.60
Residual	356660.487	3993	89.3214342	Prob > F =	0.0000
Total	488284.417	3999	122.10163	R-squared =	0.2696
				Adj R-squared =	0.2685
				Root MSE =	9.451

rwage	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
educ	2.332599	.0712024	32.76	0.000	2.193003 2.472196
exper	.1507326	.0119463	12.62	0.000	.1273111 .1741541
union	-.7856355	.4268721	-1.84	0.066	-1.622543 .0512721
female	-4.078936	.4237362	-9.63	0.000	-4.909695 -3.248176
yr2006	.3487309	.8316317	0.42	0.675	-1.281731 1.979193
yr2006female	-1.339191	.5989048	-2.24	0.025	-2.513378 -.165003
_cons	-14.10789	1.256733	-11.23	0.000	-16.57178 -11.64399

**The New York Times**

December 12, 2002, Thursday, Late Edition - Final

**Economic Scene; Sticks and stones can break bones, but the wrong name can make a job hard to find.** By

Alan B. Krueger

WHAT'S in a name? Evidently plenty if you are looking for a job.

To test whether employers discriminate against black job applicants, Marianne Bertrand of the University of Chicago and Sendhil Mullainathan of M.I.T. conducted an unusual experiment. They selected 1,300 help-wanted ads from newspapers in Boston and Chicago and submitted multiple resumes from phantom job seekers. The researchers randomly assigned the first names on the resumes, choosing from one set that is particularly common among blacks and from another that is common among whites.

So Kristen and Tamika, and Brad and Tyrone, applied for jobs from the same pool of want ads and had equivalent resumes. Nine names were selected to represent each category: black women, white women, black men and white men. Last names common to the racial group were also assigned. Four resumes were typically submitted for each job opening, drawn from a reservoir of 160. Nearly 5,000 applications were submitted from mid-2001 to mid-2002. Professors Bertrand and Mullainathan kept track of which candidates were invited for job interviews.

No single employer was sent two identical resumes, and the names on the resumes were randomly assigned, so applicants with black- and white-sounding names applied for the same set of jobs with the same set of resumes.

Apart from their names, applicants had the same experience, education and skills, so employers had no reason to distinguish among them.

The results are disturbing. Applicants with white-sounding names were 50 percent more likely to be called for interviews than were those with black-sounding names. Interviews were requested for 10.1 percent of applicants with white-sounding names and only 6.7 percent of those with black-sounding names.

Within racial groups, applications with men's or women's names were equally likely to result in calls for interviews, providing little evidence of discrimination based on sex in these entry-level jobs.

-----

Their most alarming finding is that the likelihood of being called for an interview rises sharply with an applicant's credentials -- like experience and honors -- for those with white-sounding names, but much less for those with black-sounding names. A grave concern is that this phenomenon may be damping the incentives for blacks to acquire job skills, producing a self-fulfilling prophecy that perpetuates prejudice and misallocates resources.

(Source: "Are Emily and Brendan More Employable than Lakisha and Jamal? A Field Experiment on Labor Market Discrimination" *The American Economic Review*, 2004)